

PF of Cor 8.3.11: Let $\{X_n : n \geq 0\}$ be an irreducible
aperiodic MC that has stationary distribution

$\pi = \{\pi_i\}_{i \in S}$. Then

$$\lim_{n \rightarrow \infty} P(X_n = j) = \lim_{n \rightarrow \infty} \sum_{i \in S} P(X_0 = i, X_n = j)$$

$$= \lim_{n \rightarrow \infty} \sum_{i \in S} P(X_0 = i) P(X_n = j | X_0 = i)$$

$$= \lim_{n \rightarrow \infty} \sum_{i \in S} \alpha_i P_i(X_n = j)$$

$$= \sum_{i \in S} \alpha_i \lim_{n \rightarrow \infty} P_i(X_n = j) \quad \text{by M-test (see appendix in R)}$$

$$= \underbrace{\sum_{i \in S} \alpha_i}_1 \pi_j \quad \text{by Thm 8.3.10 since } \pi \text{ is a stationary distribution}$$

$$= \pi_j$$

which gives the desired result.

Existence of Stationary Distributions (R §8.4)

Under what conditions will a stationary distribution exist?
Is it unique?

def: Given a MC_n on a state space S & given a state $i \in S$, the mean return time to state i is

$$m_i = E_i \left(\inf \{n \geq 1 : X_n = i\} \right).$$

We always have $m_i \geq 1 \quad \forall i \in S$.

If state i is transient, then $m_i = \infty$.

If state i is recurrent, then either

$m_i = \infty \leftarrow i$ is null recurrent OR

$m_i < \infty \leftarrow i$ is positive recurrent.

* Thm 8.4.1: If a Markov chain is irreducible, and if each state is positive recurrent with finite mean return time m_i , then the MC has a unique stationary distribution $\pi = \{\pi_i\}$ given by

$$\pi_i = \frac{1}{m_i} \text{ for each state } i \in S.$$

Proof of this theorem relies heavily on the following lemma.

Lemma 8.4.2: Let $G_n(i,j) = E_i[\#\{l: 1 \leq l \leq n, X_l = j\}]$
 $= \sum_{l=1}^n P_{ij}^{(l)}$. Then for an irreducible recurrent

Markov chain, $\lim_{n \rightarrow \infty} \frac{G_n(i,j)}{n} = \frac{1}{m_j}$, for any states $i \neq j$.

Pf: Let T_j^r = time of r^{th} hit of state j . Then

$$T_j^r = \underbrace{T_j^1}_{\text{i.i.d. with mean } m_j} + \underbrace{(T_j^2 - T_j^1)}_{\text{i.i.d. with mean } m_j} + \dots + \underbrace{(T_j^r - T_j^{r-1})}_{\text{i.i.d. with mean } m_j}$$

By SLLNs, 2nd version, we have that $\lim_{r \rightarrow \infty} \frac{T_j^r}{r} = m_j$ w/prob. 1. (convergence a.s.)

For $n \in \mathbb{N}$, let $r(n) = \#\{l: 1 \leq l \leq n, X_l = j\}$.

Then $\lim_{n \rightarrow \infty} r(n) = \infty$ ^{by recurrence} with prob. 1. Also,

$$T_j^{r(n)} \leq n < T_j^{r(n)+1} \quad \& \text{ now divide by } r(n)$$

$$\Rightarrow \frac{T_j^{r(n)}}{r(n)} \leq \frac{n}{r(n)} < \frac{T_j^{r(n)+1}}{r(n)} \quad \& \text{ take } \lim_{n \rightarrow \infty}$$

$$\Rightarrow \lim_{n \rightarrow \infty} \frac{n}{r(n)} = m_j \quad \text{with prob. 1}$$

$$\Rightarrow \lim_{n \rightarrow \infty} \frac{r(n)}{n} = \frac{1}{m_j} \quad \text{w/prob. 1}$$

On the other hand, $G_n(i, j) = E_i[r(n)]$ and clearly

$0 \leq \frac{r(n)}{n} \leq 1$. Then by the Bounded Convergence Thm, (Thm 7.3.1)

$$\begin{aligned} \lim_{n \rightarrow \infty} \frac{G_n(i, j)}{n} &= \lim_{n \rightarrow \infty} \frac{E_i[r(n)]}{n} = \lim_{n \rightarrow \infty} E_i \left[\frac{r(n)}{n} \right] \\ &= \frac{1}{m_j} \quad \square \end{aligned}$$

PF of Thm 8.4.1:

$$\left(E_i \left[\lim_{n \rightarrow \infty} \frac{r(n)}{n} \right] = \frac{1}{m_j} \right)$$

[Exercise for the reader!]

Prove uniqueness: $\pi_j = \frac{1}{m_j}$ is the only possible stat. dist'n.

Show that $C = \sum_j \frac{1}{m_j} = 1$.

Note: States that are not positive recurrent cannot contribute to a stationary dist'n.

Prop 8.4.5: If a MC has a stationary distribution $\pi = \{\pi_i\}$ and if a state j is not positive recurrent (i.e. $m_j = \infty$), then $\pi_j = 0$.

Cor 8.4.6: If a MC has no positive recurrent states, then it does not have a stationary dist'n.

(all $\pi_j = 0$ which contradicts $\sum_j \pi_j = 1$)

Cor 8.4.7: Let $i \neq j$ be 2 states of a MC that communicate (i.e. $f_{ij} > 0$ and $f_{ji} > 0$). If i is positive recurrent then so is j .

Cor 8.4.8: For an irreducible MC (i.e. all states communicate), either all states are positive recurrent or none are.

* Thm 8.4.9: For an irreducible MC, either

- (a) All states are positive recurrent, $\exists$ a unique stationary dist'n given by $\pi_j = \frac{1}{m_j}$ (if also aperiodic, $P_i(X_n = j) \rightarrow \pi_j$ as $n \rightarrow \infty$), OR
- (b) No states are pos. recurrent, and $\nexists$ a stationary distribution.

Example 1: Symmetric simple RW on $\mathbb{Z}$ is null recurrent, falls into category (b).

Example 2: MC on finite S necessarily falls into category (a).

Prop 8.4.10: For an irreducible MC on a finite state space, all states are positive recurrent ($\Rightarrow$ hence a unique stat. dist'n exists).

PF: Fix state $i \in S$. Let $h_{ji}^{(m)} = P_j(X_k = i \text{ for some } 1 \leq k \leq m)$
 $= \sum_{n=1}^m f_{ji}^{(n)}$.

Then $\lim_{m \rightarrow \infty} h_{ji}^{(m)} = \lim_{m \rightarrow \infty} \sum_{n=1}^m f_{ji}^{(n)} = f_{ji}$ (by def)

> 0 by irreducibility.
 (for each $j \in S$)

Since S is finite, we can find $m \in \mathbb{N}$ and $\delta > 0$ s.t. $h_{ji}^{(m)} \geq \delta \quad \forall j$.

We must also have $1 - h_{ii}^{(n)} \leq (1 - \delta)^{\lfloor n/m \rfloor}$ so that

letting $\tau_i = \inf \{n \geq 1 : X_n = i\}$, we have that

$$m_i = \sum_{n=0}^{\infty} P_i(\tau_i \geq n+1) = \sum_{n=0}^{\infty} (1 - h_{ii}^{(n)})$$

$$\leq \sum_{n=0}^{\infty} (1 - \delta)^{\lfloor n/m \rfloor}$$

$$= \frac{m}{\delta} < \infty.$$

Thus, all states are pos. recurrent.

Skip
in class

Limit Theorems

(Ref: R section 9, B section 16)

Some results needed for more advanced topics to come.

Suppose $X_1, X_2, \dots$ are random variables defined on a probability space $(\Omega, \mathcal{F}, P)$. Also, suppose that

$$\lim_{n \rightarrow \infty} X_n(\omega) = X(\omega) \quad \text{for all } \omega \in \Omega \text{ outside a set of probability 0.}$$

(for each fixed ω)

↗

In other words,
 $X_n \xrightarrow{\text{a.s.}} X$

Q. Does it follow that $\lim_{n \rightarrow \infty} E[X_n] = E[X]$?

No, not true in general.

Simple counter example:

$$\Omega = \mathbb{N}$$

$$P(\omega) = 2^{-\omega}$$

$$X_n(\omega) = \begin{cases} 2^n & \text{if } \omega = n \\ 0 & \text{if } \omega \neq n \end{cases}$$

$$\left(\begin{array}{l} \text{also denoted} \\ X_n(\omega) = 2^n \delta_{\omega,n} \end{array} \right)$$

Then $X_n \rightarrow 0$ with prob. 1 but

$$E[X_n] = 2^n \cdot 2^{-n} + 0 \cdot 2^{-\infty} = 1 \not\rightarrow 0 \text{ as } n \rightarrow \infty$$