

## Further Classification of states (§4.3 con't)

$f_i$  = Probability the process will ever re-enter state  $i$ , starting in state  $i$

def: State  $i$  is recurrent if  $f_i = 1$ .

The process will re-enter state  $i$  infinitely often

def: State  $i$  is transient if  $f_i < 1$ .

Let  $X$  = # of times the process is in state  $i$ .  
(starting from  $i$ )

$$P(X=n) = f_i^{n-1} \overset{\text{"success" prob.}}{(1-f_i)}, \quad n=1, 2, 3, \dots$$

↑  
geometric distribution with mean

$$E[X] = \frac{1}{1-f_i} < \infty$$

(alternate version of geometric)

$$P(X=n) = f_i^n (1-f_i), \quad n=0, 1, 2, \dots$$

(Prop. 4.1)

Fact: state  $i$  is recurrent iff  $\sum_{n=1}^{\infty} P_{ii}^n = \infty$ .

State  $i$  is transient iff  $\sum_{n=1}^{\infty} P_{ii}^n < \infty$ .

A transient state will only be visited a finite # of times.

Thus, for a finite state Markov chain, not all states can be transient. Why not?

More details ... on Prop 4.1

\* state  $i$  recurrent iff starting in  $i$ , the expected # times process is in  $i$  is infinite.

$$\text{Let } I_n = \begin{cases} 1 & \text{if } X_n = i \\ 0 & \text{if } X_n \neq i \end{cases}$$

Then  $\sum_{n=0}^{\infty} I_n = \#$  time periods the process is in state  $i$

$$E \left[ \sum_{n=0}^{\infty} I_n \mid X_0 = i \right] = \sum_{n=0}^{\infty} E[I_n \mid X_0 = i] \quad \text{linearity of Exp. Value}$$

$$= \sum_{n=0}^{\infty} P(X_n = i \mid X_0 = i) \cdot 1 + P(X_n \neq i \mid X_0 = i) \cdot 0$$

$$= \sum_{n=0}^{\infty} P_{ii}^n$$

$$\text{So, } i \text{ recurrent} \iff \sum_{n=0}^{\infty} P_{ii}^n = \infty$$

$$i \text{ transient} \iff \sum_{n=0}^{\infty} P_{ii}^n < \infty$$

Fact: Recurrence & Transience are class properties.

i.e. All states in 1 equivalence class are either all recurrent OR all transient.

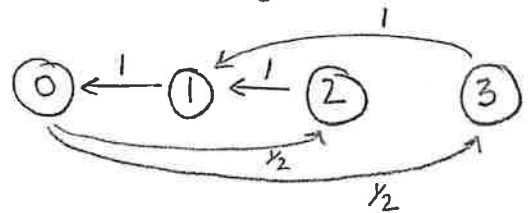
Note: If a Markov chain is finite and irreducible (meaning all states belong to the same class), then all states are recurrent.

\*R  
MCSim3  
Ex 1

Example: state space =  $\{0, 1, 2, 3\}$

$$P = \begin{bmatrix} 0 & 0 & \frac{1}{2} & \frac{1}{2} \\ 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 1 & 0 & 0 \end{bmatrix}$$

Draw Diagram:



Q. Which states are transient?  
Which states are recurrent?

Check your work in R.

↖ All states are recurrent  
⇒ irreducible MC

## A few remarks:

- If  $i$  is recurrent &  $j$  is transient, then  $i \not\leftrightarrow j$ .
- Starting from a recurrent state, the MC cannot enter a transient state.  
(Converse is possible)
- Once a MC enters a recurrent state, it can never leave that class.

§4.4

For pairs of states  $i \neq j$ , let

$$f_{ij} = P(X_n = j \text{ for some } n > 0 \mid X_0 = i)$$

prob. that the MC, starting in state  $i$ , will ever make a transition into state  $j$

If state  $i$  is recurrent and  $i \leftrightarrow j$ , then

$$f_{ij} = 1$$

Q. What is the average time<sup>(# of transitions needed)</sup> to come back to  $i$ ? when starting in  $i$

$$T_i = \min \{ n \geq 1 : X_n = i \} \quad \text{— first time returning to } i$$

$$P(T_i = n) = f_i^{(n)} \quad \text{— prob. first return takes } n \text{ steps}$$

$$\text{mean time: } E[T_i] = \sum_{n=1}^{\infty} n f_i^{(n)} = m_i \quad \text{(same notation as book)}$$

See pg. 30 instead

Long-run Proportions § 4.4

Let  $i$  be a recurrent state of DTMC  $\{X_n : n=0,1,\dots\}$ .

Let  $m_i$  = expected # of transitions that MC takes, when starting in state  $i$ , to return to  $i$ .

$$m_i = E[N_i | X_0 = i] \quad \text{where} \quad N_i = \min\{n > 0 : X_n = i\}$$

$\uparrow$  expected # transitions                       $\uparrow$  # of transitions

def: The recurrent state  $i$  is

- positive recurrent if  $m_i < \infty$  and
- null recurrent if  $m_i = \infty$

Now suppose the MC is irreducible & recurrent. Then the long-run proportion of time the chain spends in state  $i$  is  $\frac{1}{m_i}$ .

Let  $\pi_i$  denote the long-run proportion of time spent in state  $i$ , then for any initial state,

$$\boxed{\pi_i = \frac{1}{m_i}} \quad \forall i \in S.$$

Proof: See Models § 4.4 Prop 4.4