

Recap of Discrete-time Markov Chains (DTMC)

$\{X_n : n=0, 1, \dots\}$ is a DTMC on state space S .

Markov Property:

$$P(X_{n+1} = j | X_n = i) = P_{ij} \leftarrow \text{entry in } P \text{ (transition matrix)}$$

is independent of past history of MC, only depends on current state $X_n = i$.

Classification of States

$i \leftrightarrow j$ $i \neq j$ communicate

If all states communicate, MC is irreducible
recurrent vs transient states

$\begin{cases} \downarrow \\ \text{positive recurrent if } m_i < \infty \\ \rightarrow \\ \text{null recurrent if } m_i = \infty \end{cases}$

where $m_i = E[N_i | X_0 = i]$, $N_i = \min\{n \geq 0 : X_n = i\}$

$\begin{array}{l} \swarrow \text{expected \# transitions} \\ \text{to return to } i \end{array} \quad \begin{array}{l} \searrow \# \text{ of transitions} \\ \text{until MC moves} \\ \text{to state } i \end{array}$

Note: Quantity $\sum_{n=0}^{\infty} I_n$ counts # of times MC is in state i

NOT the same as m_i !!

(e.g. Say MC visits state i 10 times by time n , exp. return time is 50 time steps.)

- If i transient:

$$\sum_{n=1}^{\infty} P_{ii}^n < \infty \quad \begin{matrix} \text{— only visit transient} \\ \text{state a finite # of times} \end{matrix}$$

$$P(N_i = \infty | X_0 = i) > 0 \quad \text{and} \quad m_i = \infty$$

↑
exp. return time

- If i is recurrent:

$$\sum_{n=1}^{\infty} P_{ii}^n = \infty \quad \begin{matrix} \text{— return to recurrent} \\ \text{state infinitely often} \end{matrix}$$

↙ If $m_i < \infty$, then i is positive recurrent

↙ If $m_i = \infty$, then i is null recurrent

Interpretation

A positive recurrent state - returns are relatively quick.

A null recurrent state - can take a very long time for MC to return to that state.
(mean time = ∞)

Note: If a MC is finite, all recurrent states are positive recurrent.

(Can only have null recurrent state if S has ∞ # of states!)

Limiting Probabilities, Stationary Distribution § 4.4 (con't)

Let $\{\pi_j, j \geq 0\}$ be the long-run proportion of transitions that come from state j

Then $\pi_i P_{ij} = \text{long-run prop. of transitions from } i \text{ to } j$

Thm: Consider an irreducible MC. If the chain is positive recurrent, then the long-run proportions are the unique solution of

$$\begin{aligned} \textcircled{*} \quad \left\{ \begin{array}{l} \pi_j = \sum_{i \in S} \pi_i P_{ij}, \quad j \geq 0 \\ \sum_{j \in S} \pi_j = 1. \end{array} \right. \end{aligned}$$

If there is no solution to eqns above, then MC is either transient or null recurrent and all $\pi_j = 0$.

→ see EX →

Note: The long-run proportions $\pi_j, j \geq 0$ are called stationary probabilities.

$\pi = [\pi_1, \pi_2, \dots, \pi_j, \dots]$ — stationary distribution
(for all $j \in S$)

More compactly written: $\boxed{\pi = \pi P}$
vector matrix

Main Idea:

If $P(X_0 = j) = \pi_j$, then $P(X_n = j) = \pi_j \quad \forall n, j \geq 0$

start in state j /
 be in state j
 after n steps

→ Example: Simple random walk on \mathbb{Z}

- symmetric RW is null recurrent ($p = \frac{1}{2}$)
- asymmetric RW is transient ($p \neq \frac{1}{2}$)

Limiting Probabilities

Consider the 2-state DTMC given by

$$P = \begin{matrix} & 0 & 1 \\ 0 & 0.7 & 0.3 \\ 1 & 0.4 & 0.6 \end{matrix}$$

Note that

$$P^4 = \begin{bmatrix} 0.575 & 0.425 \\ 0.564 & 0.433 \end{bmatrix} \quad \ddots \quad P^8 = \begin{bmatrix} 0.572 & 0.428 \\ 0.570 & 0.430 \end{bmatrix}$$

Very close
rows are almost identical

It seems that the long-run proportions

$$\pi_0 = \frac{4}{7} \approx 0.571 \quad \& \quad \pi_1 = \frac{3}{7} \approx 0.429$$

(computed in book § 4.4 Ex. 4.20) → solve eqns \star
to get these

may also be the limiting probabilities.

True in this case, but NOT true in general.

Counterexample - Periodic Chain

$$P = \begin{bmatrix} 0 & 1 \\ 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \text{Diagram: } \textcircled{0} \xrightarrow{1} \textcircled{1} \quad \text{with } 1 \text{ below the arrow}$$

This MC continually alternates b/t states 0 & 1
so $\pi_0 = \pi_1 = \frac{1}{2}$.

BUT $P_{00}^n = \begin{cases} 1 & \text{if } n \text{ even} \\ 0 & \text{if } n \text{ odd} \end{cases}$ so $P_{00}^n \rightarrow ??$
as $n \rightarrow \infty$

does NOT have a
limiting value!!

* If a MC can only return to a state in a multiple
of $d > 1$ steps ($d=2$ in above example), it is
called periodic; does NOT have limiting probabilities.

def: A MC that is not periodic is called aperiodic,
(all states have period 1.)

Thm: For an irreducible MC that is aperiodic,
limiting probabilities always exist \Rightarrow do NOT
depend on the initial state i .

$$\lim_{n \rightarrow \infty} P_{ij}^n \text{ exists}$$

Further, $\pi_j^* = \lim_{N \rightarrow \infty} \underbrace{\frac{\alpha_j(N)}{N}}_{\substack{\nearrow \\ \text{limiting prob.}}}$

For large N , this is
long-run proportion of
time described before (π_j^*)

π_j^* satisfies eqns \circledast ,
unique soln.

where

$\alpha_j(N) = \text{amount of}$
time MC spends
in state j
during time
 $1, 2, \dots, N$

def: State i has period d if

$P_{ii}^n = 0$ whenever n is NOT divisible by d ,
and d is the largest integer with this
property.

Periodicity = Class property : if $i \leftrightarrow j$, then $i \& j$
have same period.

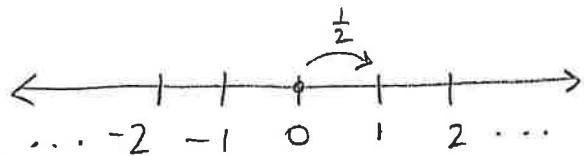
def: A aperiodic state is a state with period 1.

Periodicity = Class property: If $i \leftrightarrow j$, then $i \not\sim j$ have the same period

Example: Simple RW on \mathbb{Z}

Lec 11 (3b)

- If you leave a state, can only come back in an even # of steps.
- Period of any state is 2. (Verify in R.)



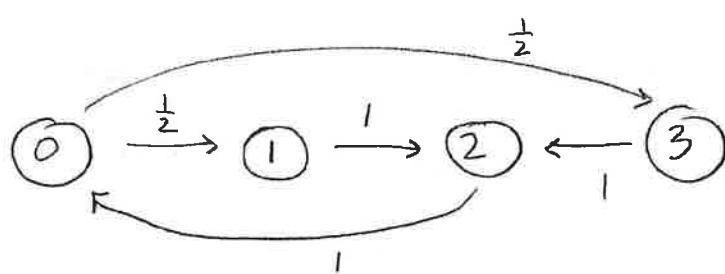
e.g. $0 \rightarrow 1 \rightarrow 0$ Return to 0 in 2 steps

Example: MC on 4 states with

did last Thurs

$$P = \begin{bmatrix} 0 & \frac{1}{2} & 0 & \frac{1}{2} \\ 0 & 0 & 1 & 0 \\ 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 \end{bmatrix}$$

All states have period 3.
(Verify this in R.)



e.g. $0 \rightarrow 1 \rightarrow 2 \rightarrow 0$
OR

$0 \rightarrow 3 \rightarrow 2 \rightarrow 0$

Example 1: 3 state MC with transition matrix

$$P = \begin{bmatrix} 0.5 & 0.4 & 0.1 \\ 0.3 & 0.4 & 0.3 \\ 0.2 & 0.3 & 0.5 \end{bmatrix}$$

In the long run, what proportion of time is the process in each of the 3 states?

Soln: ① Solve $\pi = \pi P$ (i.e. $\pi_j = \sum_{i=0}^2 \pi_i P_{ij} \quad \forall j$)

$$\left\{ \begin{array}{l} \pi_0 = 0.5 \pi_0 + 0.3 \pi_1 + 0.2 \pi_2 \\ \pi_1 = 0.4 \pi_0 + 0.4 \pi_1 + 0.3 \pi_2 \\ \pi_2 = \dots \end{array} \right. \quad \text{subject to } \pi_0 + \pi_1 + \pi_2 = 1$$

OR ② put P into $R \in \mathbb{R}^{3 \times 3}$ raise it to a large enough power to see columns converge to

~~Dⁱⁿ~~

$$P_i^n = [\pi_0 \quad \pi_1 \quad \pi_2]$$

↑
row i
(all same
after some n)

Ans: $\pi_0 = \frac{21}{62} \approx 0.3387, \pi_1 = \frac{23}{62} \approx 0.3710, \pi_2 = \frac{18}{62} \approx 0.2903$

Recall example from last week: $\{0,1\}$, $\{3\}$, $\{2\}$

no unique stationary distribution

recurrent transient
 /

Mean Time Spent in Transient States § 4.6

SKIP Details

Recall: Once a MC enters a recurrent state, it can never leave that class

⇒ Each recurrent class is a closed class

state i recurrent & $i \not\leftrightarrow j$ (does NOT communicate),
then $P_{ij} = 0$.

SKIP

In particular,

$$\begin{array}{l} k \text{ recurrent} \\ j \text{ transient} \end{array} \} \Rightarrow P_{kj} = 0$$

Consider a finite state MC.

$T = \{1, 2, \dots, t\}$ denotes the set of transient states

$$P_T = \begin{bmatrix} P_{11} & \cdots & P_{1t} \\ P_{21} & \cdots & P_{2t} \\ \vdots & & \vdots \\ P_{t1} & \cdots & P_{tt} \end{bmatrix} \quad \text{matrix of 1-step transition probabilities from transient states to transient states}$$

just say

(* sum of rows might be less than 1 ! why ?
 o.w. T would be a closed class of states)

For $i, j \in T$,

s_{ij} = expected # of times MC is in state j given that it starts in state i

$$\Rightarrow s_{ij} = \delta_{ij} + \sum_{k=1}^t p_{ik} s_{kj}$$

Kronecker delta

$$\delta_{ij} = \begin{cases} 1 & \text{if } i=j \\ 0 & \text{if } i \neq j \end{cases}$$

In Matrix Notation : $S = I + P_T S$

where

$$S = \begin{bmatrix} s_{11} & \cdots & s_{1t} \\ \vdots & \ddots & \vdots \\ s_{t1} & \cdots & s_{tt} \end{bmatrix}$$

Identity matrix ($t \times t$)

Solving for S yields:

$$\underline{S} = (I - P_T)^{-1}$$

Details

$$S = I + P_T S$$

$$S - P_T S = I$$

$$S(I - P_T) = I$$

$$S = (I - P_T)^{-1}$$

Say * Quantities s_{ij} , $i, j \in T$ can be obtained by inverting the matrix $I - P_T$

Example (to do in R) : Gambler's Ruin

$$p = 0.4$$

$$N = 7$$

$$P_T = \begin{pmatrix} 1 & 2 & 3 & 4 & 5 & 6 \\ 1 & 0 & 0.4 & & & \\ 2 & 0.6 & 0 & 0.4 & & 0 \\ 3 & & 0.6 & 0 & 0.4 & \\ 4 & & & 0.6 & 0 & 0.4 \\ 5 & & 0 & & 0.6 & 0 & 0.4 \\ 6 & & & & & 0.6 & 0 \end{pmatrix}$$

Starting with \$3,

- compute the expected amount of time the gambler has \$5
 - " " " gambler has \$2

i.e. Find $s_{3,5}$ and $s_{3,2}$

* use solve function
in R

For $i, j \in T$,

f_{ij} = probability MC ever transitions into j
given that it starts in i



can be computed
from s_{ij} 's

Key Relation

$$s_{ij} = s_{ij} + f_{ij} s_{jj}$$

$$\Rightarrow f_{ij} = \frac{s_{ij} - s_{ij}}{s_{jj}}$$