

Markov Chains

[ref: Models ch. 4]

Discrete-Time
Markov Chain

vs.

Continuous-Time
Markov Chain

index set $\rightarrow T = \mathbb{N} = \{0, 1, 2, \dots\}$
(or finite subset)

 $T = [0, \infty)$ State space = \mathbb{N} \leftarrow for examplestate space = \mathbb{N}

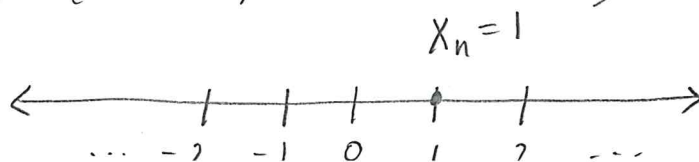
* Markov chain is a special type of stochastic process which has the Markov Property:

the conditional distribution of the future state given the current state & the past is independent of the past.

In other words,

the probability of any particular behavior of the process, when the current state is known, is not affected by additional knowledge of its past behavior.

e.g. Random Walk (1-dim)
on \mathbb{Z}

 $\rightarrow X_{n+1} = 0 \text{ or } 2$

doesn't care
about X_i for
 $i = 0, 1, \dots, n-1$

$\{X_n : n=0,1,\dots\}$ - discrete-time MC

Transition Probability from state i to state j at time n

Discrete time case

$$P(X_{n+1} = j \mid X_n = i, \underbrace{X_{n-1} = i_{n-1}, \dots, X_0 = i_0}_{\text{past history}})$$

(1-step transition)

$$= P(\underbrace{X_{n+1} = j}_{\text{future state}} \mid \underbrace{X_n = i}_{\text{current state}})$$

where $i_0, i_1, \dots, i_{n-1}, i, j, n \geq 0$

$$= P_{ij}(n)$$

* Independent of past history!

Continuous time case

$$P(X_{t+s} = j \mid X_s = i, X_u = x(u), 0 \leq u < s)$$

all times u less than current time s

$$= P(X_{t+s} = j \mid X_s = i)$$

where $i, j, x(u) \geq 0$
 $0 \leq u < s$

$$= P_{ij}(t)$$

In many examples, the transition probability does not depend on the time n :

$$\text{If } \boxed{P(X_{n+1} = j \mid X_n = i) = P_{ij}} = P(X_1 = j \mid X_0 = i)$$

then the Markov chain has

stationary transition probabilities.

↖ No dependence on n !

Arrange the P_{ij} 's into a Transition Probability Matrix P
 i.e. matrix of 1-step transition probabilities

$$P = \begin{pmatrix} P_{00} & P_{01} & P_{02} & \dots \\ P_{10} & P_{11} & P_{12} & \dots \\ P_{20} & P_{21} & P_{22} & \dots \\ \vdots & \vdots & \vdots & \vdots \\ P_{i0} & P_{i1} & P_{i2} & \dots \\ \vdots & \vdots & \vdots & \vdots \end{pmatrix}$$

$$P_{ij} = P(X_{n+1} = j \mid X_n = i)$$

= prob. that the process makes the transition $i \rightarrow j$

$$P_{ij} \geq 0 \quad \& \quad \sum_{j=0}^{\infty} P_{ij} = 1 \quad \text{— each row of } P \text{ sums to } 1$$

for $i=0,1,2,\dots$

Since the process must make a transition into some state each time step

→ First Do Examples!

n -step transition probabilities

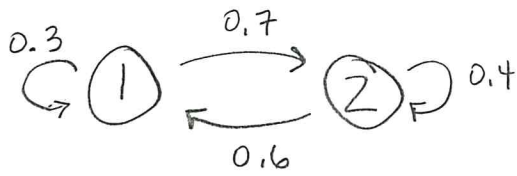
$$P_{ij}^n = P(X_{n+k} = j \mid X_k = i), \quad n \geq 0, \quad i, j \geq 0$$

\ process moves from state i to j
 in n steps (transitions)

Q. Long term behavior of the MR?

Maybe Lec 9 ??

Example 1: 2-state Markov chain



$$P = \begin{matrix} & \begin{matrix} 1 & 2 \end{matrix} \\ \begin{matrix} 1 \\ 2 \end{matrix} & \begin{bmatrix} 0.3 & 0.7 \\ 0.6 & 0.4 \end{bmatrix} \end{matrix}$$

move to next state in 1 step

rows sum to 1

current state

Example 2: Wright-Fisher model (genetics)

Model of random genetic drift
change in allele frequency of a gene over time

Consider a constant size population of $2N$ individuals:
2 types $\begin{cases} \text{A allele} \\ \text{a allele} \end{cases}$

If the current generation consists of
 $\begin{cases} i \text{ type A individuals} \\ 2N-i \text{ type a individuals} \end{cases}$

then the next generation is determined by $2N$ Bernoulli trials with "success" probability

$$P_i = \frac{i}{2N} \quad (\text{for an A type indiv.})$$

$$\text{and } \left(1 - P_i = 1 - \frac{i}{2N} = \frac{2N-i}{2N} \quad (\text{for an a type indiv.}) \right)$$

Define a MC $\{X_n : n=0,1,\dots\}$ such that

$X_n = \#$ of type A alleles in the n^{th} generation

$$S = \{0, 1, \dots, 2N\}$$

i.e. Binomial RV
w/parameters
 $n \hat{=} 2N$
 $p \hat{=} p_i$

Transition Probability:

$$P(X_{n+1} = j \mid X_n = i) = \underbrace{\binom{2N}{j} p_i^j (1-p_i)^{2N-j}}_{\text{Binomial PDF}} \quad \text{for } i, j = 0, 1, \dots, 2N$$

Note: $0 \hat{=} 2N$ are called absorbing states

i.e. once the process enters 1 of these states
(where all individuals are the same type),
the process cannot leave that state.

If we modify the model to include mutation:

$$\mu_1 = P(A \rightarrow a) \quad \hat{=} \quad \mu_2 = P(a \rightarrow A)$$

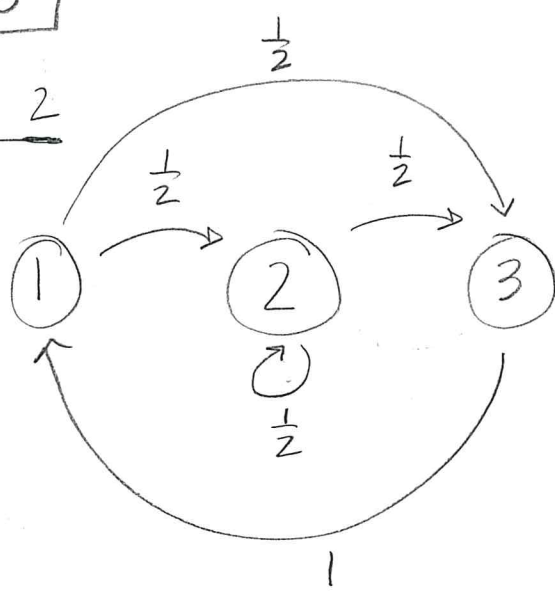
$$\text{then } p_i = \frac{i}{2N} (1 - \mu_1) + \left(1 - \frac{i}{2N}\right) \mu_2$$

$\Rightarrow 0 \hat{=} 2N$ are no longer absorbing states

$\hat{=} \text{ long-term behavior of MC}$
stabilizes as $n \rightarrow \infty$.

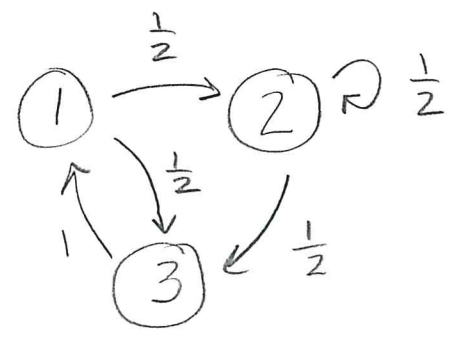
Lecture 8

Example 2



R Practice

9/21/17 (4)
Stat 753

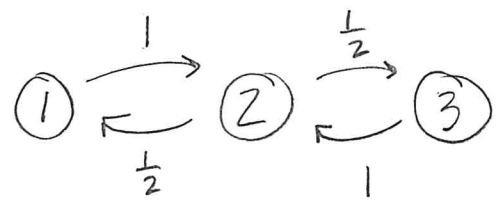


$$P = \begin{matrix} & \begin{matrix} 1 & 2 & 3 \end{matrix} \\ \begin{matrix} 1 \\ 2 \\ 3 \end{matrix} \text{ states} & \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ 0 & \frac{1}{2} & \frac{1}{2} \\ 1 & 0 & 0 \end{bmatrix} \end{matrix}$$

Interpret as

Prob $1 \rightarrow 2 = \frac{1}{2}$
row col

Example 1



$$P = \begin{matrix} & \begin{matrix} 1 & 2 & 3 \end{matrix} \\ \begin{matrix} 1 \\ 2 \\ 3 \end{matrix} \text{ states} & \begin{bmatrix} 0 & 1 & 0 \\ \frac{1}{2} & 0 & \frac{1}{2} \\ 0 & 1 & 0 \end{bmatrix} \end{matrix}$$

Check:

* Rows of P sum to 1 ✓

Example 3

-modified version of Example 2:

2 is an absorbing state

$$P = \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ 0 & 1 & 0 \\ 1 & 0 & 0 \end{bmatrix}$$